

# The Nubeam reference-free approach to analyze metagenomic sequencing reads (Supplemental information)

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## 1 Supplemental Material and Methods

### 1.1 Simulation of genomes and inferences

We simulated genome for each taxonomic unit in the designed trees, with the root genome of *E. coli* strain K-12 substrain MG1655 (4,641,652 bp). We used iSG (Strope et al., 2006) for simulation under the general time-reversible (GTR) substitution model. The 6 parameters for relative substitution rate are as those estimated from the evolution study of bacteria *rrsA*, and 4 parameters for nucleotide frequency are those estimated from the *E. coli* genome. The branch lengths represent number of substitutions per basepair per generation. For the three-taxonomic-unit trees, both branch lengths of S1 and S2 had fixed values of  $1 \times 10^{-4}$ , while the branch length of S3 had values span two orders of magnitude from  $10^{-2}$  to  $10^{-4}$ . For the eight-taxonomic-unit trees, we used six combinations of internal and terminal branch lengths of  $1 \times 10^{-5}$  and  $5 \times 10^{-5}$  to represent different difficulty levels of phylogeny inference.

Each simulation was done 100 replicates. For three-taxonomic-tree we only interested in relative distance between taxa. For eight-taxonomic-unit tree, we used hierarchical clustering (with Wards minimum variance method) to reconstruct trees from distance matrices produced by both Nubeam and the *k*-mer method. We compared the reconstructed trees with true trees by Compare2Trees (Nye et al., 2005), generating a score ranging from 0 to 1 for each comparison, with 0 indicating no topological similarity and 1 indicating same topology.

### 1.2 Bin partitioning

The following algorithm does the balanced partition of bins.

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**Algorithm 1** Balanced partition of bins using conditioning quantiles

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1: Combine the matrices for all samples
2: Partition 1st column into BINs using quantiles
3: for each BIN of 1st dimension do
4:   for each ROW do
5:     collect the ROW if its 1st NUMBER is in BIN
6:   end for
7:   for the collected rows, partition their 2nd column into BINs using quantiles
8:   for each BIN of 2nd dimension do
9:     for each ROW of the collected rows do
10:      collect the ROW if its 2nd NUMBER is in BIN
11:    end for
12:    for the collected rows, partition their 3rd column into BINs using quantiles
13:    ...
14:    for each BIN of  $(m - 1)$ th dimension do
15:      for each ROW of the collected rows do
16:        collect the ROW if its  $(m - 1)$ th NUMBER is in BIN
17:      end for
18:      for the collected rows, partition their  $m$ -th column into BINs using quantiles
19:    end for
20:  end for
21: end for
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### 1.3 $k$ -mer frequency based method is a special case of Nubeam

In  $k$ -mer frequency based methods, empirical probability distributions of  $k$ -mers are generated for the sequencing samples to be compared. For a determined  $k$ , consider set  $\mathcal{K}$  as the set of all possible  $k$ -mers determined by the four nucleotides. The number of all possible  $k$ -mers  $|\mathcal{K}|$  is  $4^k$ . For  $i_{th}$  sequencing sample with  $n$  collected  $k$ -mers, let  $N_w^{(i)}$  be the frequency of a given  $k$ -mer  $w$ . The relative frequency, or empirical probability, of  $w$  in the sample is  $f_w^{(i)} = \hat{\mathbb{P}}_i(w) = \frac{N_w^{(i)}}{\sum_{r \in \mathcal{K}} N_r^{(i)}} = \frac{N_w^{(i)}}{n}$ . The empirical probability distributions of  $k$ -mers in  $i_{th}$  and  $j_{th}$  sequencing sample are:

$$\hat{\mathbb{P}}_i = \begin{bmatrix} f_{w_1}^{(i)} & f_{w_2}^{(i)} & f_{w_3}^{(i)} & \dots & f_{w_{|\mathcal{K}|}}^{(i)} \end{bmatrix}$$

$$\hat{\mathbb{P}}_j = \begin{bmatrix} f_{w_1}^{(j)} & f_{w_2}^{(j)} & f_{w_3}^{(j)} & \dots & f_{w_{|\mathcal{K}|}}^{(j)} \end{bmatrix}$$

Note that each  $k$ -mer is treated as a single category, so  $f_w$  do not need to be ordered in a specific way, but to compare the two distributions  $f_w$  must be ordered in a same way for two samples. If a  $k$ -mer  $w$  is absent from a sequencing sample, then  $f_w = 0$ . Next the two distributions are compared using distance or dissimilarity measures.

In Nubeam, each unique  $k$ -mer  $w$  is represented by an unique vector of numbers  $\mathbf{y}_w = [y_{1w} \ y_{2w} \ y_{3w} \ y_{4w}]$ ,  $\mathbf{y}_w \in \mathbb{R}^4$ . The  $i_{th}$  sequencing sample with  $n$  collected  $k$ -mers is represented by an  $n \times 4$  matrix:

$$\mathbf{y}_i = [\mathbf{y}_{i1} \ \mathbf{y}_{i2} \ \mathbf{y}_{i3} \ \mathbf{y}_{i4}] = \begin{bmatrix} y_{i11} & y_{i21} & y_{i31} & y_{i41} \\ y_{i12} & y_{i22} & y_{i32} & y_{i42} \\ \vdots & \vdots & \vdots & \vdots \\ y_{i1n} & y_{i2n} & y_{i3n} & y_{i4n} \end{bmatrix}$$

The  $\mathbf{y}_i$  could be viewed as an 4-dimensional multivariate distribution of random variables  $Y_{i1}, Y_{i2}, Y_{i3}, Y_{i4}$ , with probability distribution function (pdf) of  $f_{Y_{i1}, Y_{i2}, Y_{i3}, Y_{i4}}(y_{i1}, y_{i2}, y_{i3}, y_{i4})$ . Here each  $\mathbf{y}_w$  is a data point and has a mass of  $\frac{1}{n}$ . Divide  $\mathbb{R}^4$  into a total of  $b$  disjoint sets, i.e., a total of  $b$  4-dimensional hyper-rectangle bins  $\mathcal{B}_1, \mathcal{B}_2, \dots, \mathcal{B}_b$ . Define the empirical probability distribution by

$$\hat{\mathbb{P}}_i = \begin{bmatrix} f_{\mathcal{B}_1}^{(i)} & f_{\mathcal{B}_2}^{(i)} & f_{\mathcal{B}_3}^{(i)} & \dots & f_{\mathcal{B}_b}^{(i)} \end{bmatrix}$$

with  $f_{\mathcal{B}} = \hat{\mathbb{P}}(\mathcal{B}) = \frac{\sum_{w=1}^n I_{\mathcal{B}}(\mathbf{y}_w)}{n}$ . Next we quantify the genetic distance between the  $i_{th}$  sample and  $j_{th}$  sample by comparing the two empirical probability distributions using distance or dissimilarity measures.

When binning of  $\mathbb{R}^4$  is fine enough such that  $\max_{\mathcal{B} \in \{\mathcal{B}_1, \mathcal{B}_2, \dots, \mathcal{B}_b\}} |\mathcal{B}| = 1$ , Nubeam would be identical with  $k$ -mer frequency based method.

PROOF. We need to prove that for each of the samples to be compared, the vectors representing the empirical probability distributions in  $k$ -mer frequency based method

$\hat{\mathbb{P}}_{k\text{-mer}} = [f_{w_1} \ f_{w_2} \ f_{w_3} \ \dots \ f_{w_{|\mathcal{K}|}}]$  and Nubeam  $\hat{\mathbb{P}}_{Nubeam} = [f_{\mathcal{B}_1} \ f_{\mathcal{B}_2} \ f_{\mathcal{B}_3} \ \dots \ f_{\mathcal{B}_b}]$  are equivalent.

Each unique  $k$ -mer  $w$  is represented by an unique vector of numbers  $\mathbf{y}_w$ . For any specific  $k$ -mer  $r$ , suppose  $\mathbf{y}_r \in \mathcal{B}_r$ . Since

$$\max_{\forall \mathcal{B} \in \{\mathcal{B}_1, \mathcal{B}_2, \dots, \mathcal{B}_b\}} |\mathcal{B}| = 1$$

we must have  $|\mathcal{B}_r| = 1$ . Hence,

$$\begin{aligned} \frac{\sum_{w=1}^n I_{\mathcal{B}_r}(\mathbf{y}_w)}{n} &= \frac{N_r}{n} \\ \Rightarrow \hat{\mathbb{P}}(\mathcal{B}_r) &= \hat{\mathbb{P}}(r) \\ \Rightarrow f_{\mathcal{B}_r} &= f_r \end{aligned}$$

For  $k$ -mer frequency based method, in  $\hat{\mathbb{P}}_{k\text{-mer}} = [f_{w_1} \ f_{w_2} \ f_{w_3} \ \dots \ f_{w_{|\mathcal{K}|}}]$ ,  $f_w$  do not need to be ordered in a specific way;  $f_w$  only need to be ordered in a same way for two samples to be compared. In Nubeam, to meet this requirement, for each of the samples to be compared, we only need to divide  $\mathbb{R}^m$  into bins in a same way and order these bins in a same way. In practice, if the binning is a function of data, we can combine the samples to be compared and then do the binning. Hence, for each of the samples to be compared,  $\hat{\mathbb{P}}_{k\text{-mer}} \equiv \hat{\mathbb{P}}_{Nubeam}$ .  $\square$

Thus,  $k$ -mer frequency based method is a special case of Nubeam, which is a more general idea.



## 2 Supplemental Figures

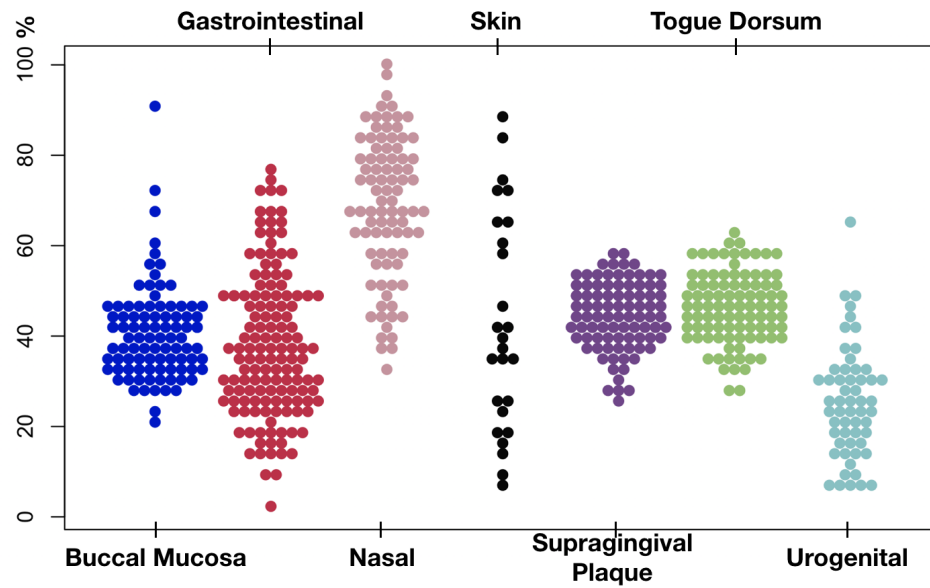


Figure S1: Percent of unmapped reads by body sites. Nasal samples have highest average unmapped reads, and gastrointestinal and urogenital samples have the lowest. Skin samples have large spread of unmapped reads.

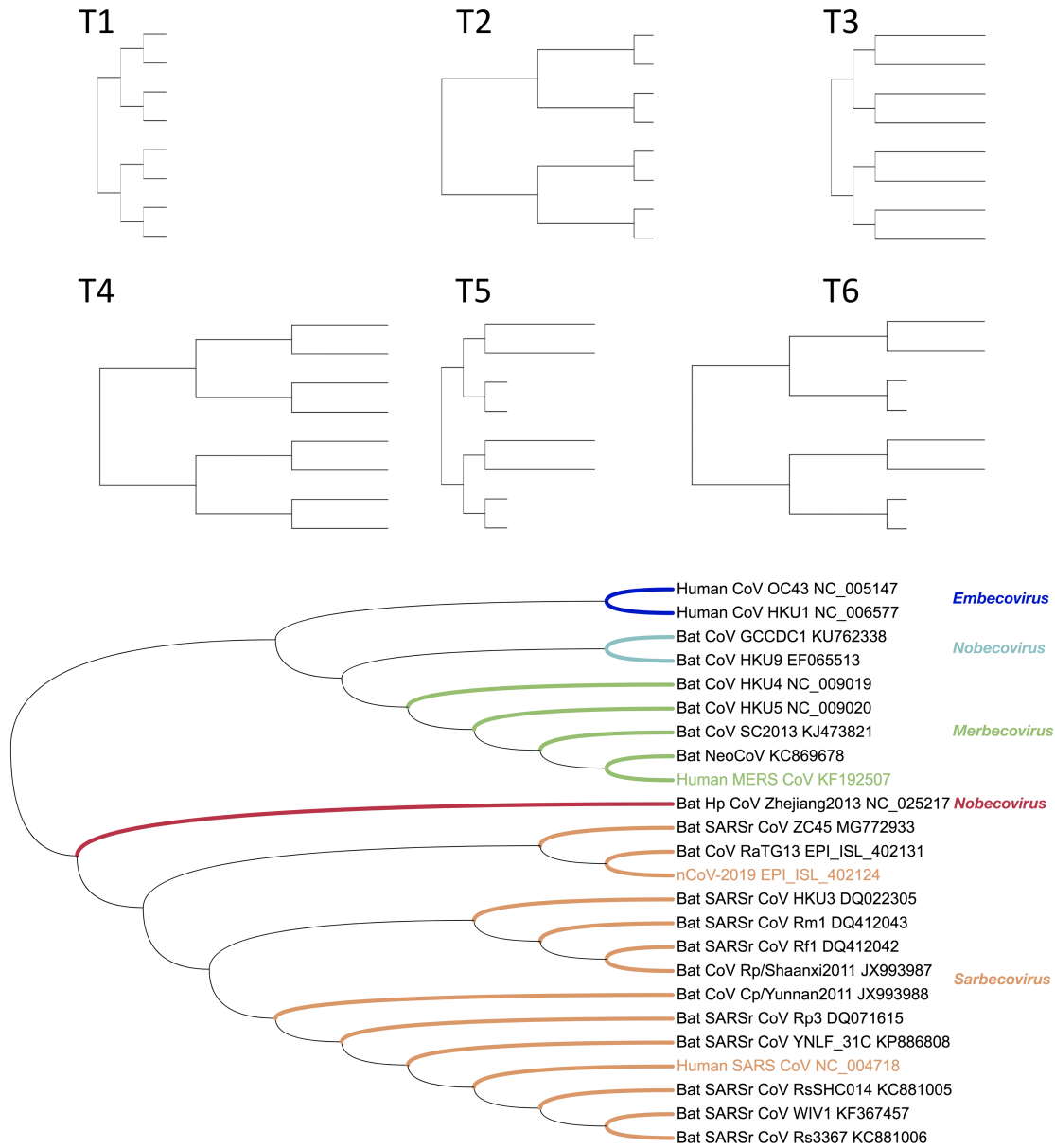


Figure S2: Top: Six combinations of internal and terminal branch lengths of 0.0001 and 0.0005 were used to represent different degrees of genome divergence and thus different difficulty levels of phylogeny inference. Bottom: Nubeam clustered the beta-coronavirus complete genomes according to sub-genera, correctly grouped human SARS and MERS coronavirus with closest bat strains at the genome level and recapitulated the findings that the genome of bat coronavirus RaTG13 is highly similar to that of nCoV-2019. The clustering was presented using iTOL (Letunic and Bork, 2019).

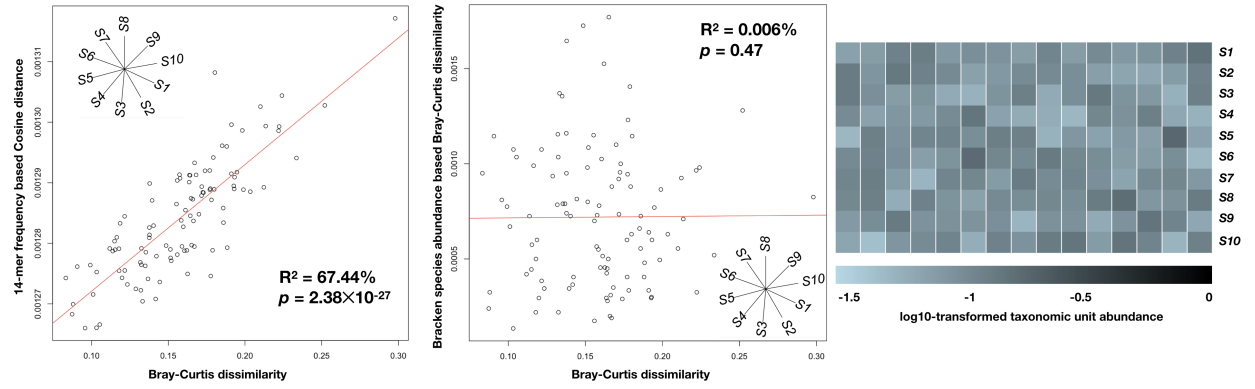


Figure S3: The dissimilarities based on 14-mer frequency Lu et al. (2017b) and Bracken-estimated species abundance Lu et al. (2017a) have weaker linear relationship with composition-based dissimilarities among synthetic communities. The significance of linear relationship is measured by  $R^2$  and  $p$ -value for regression coefficient.

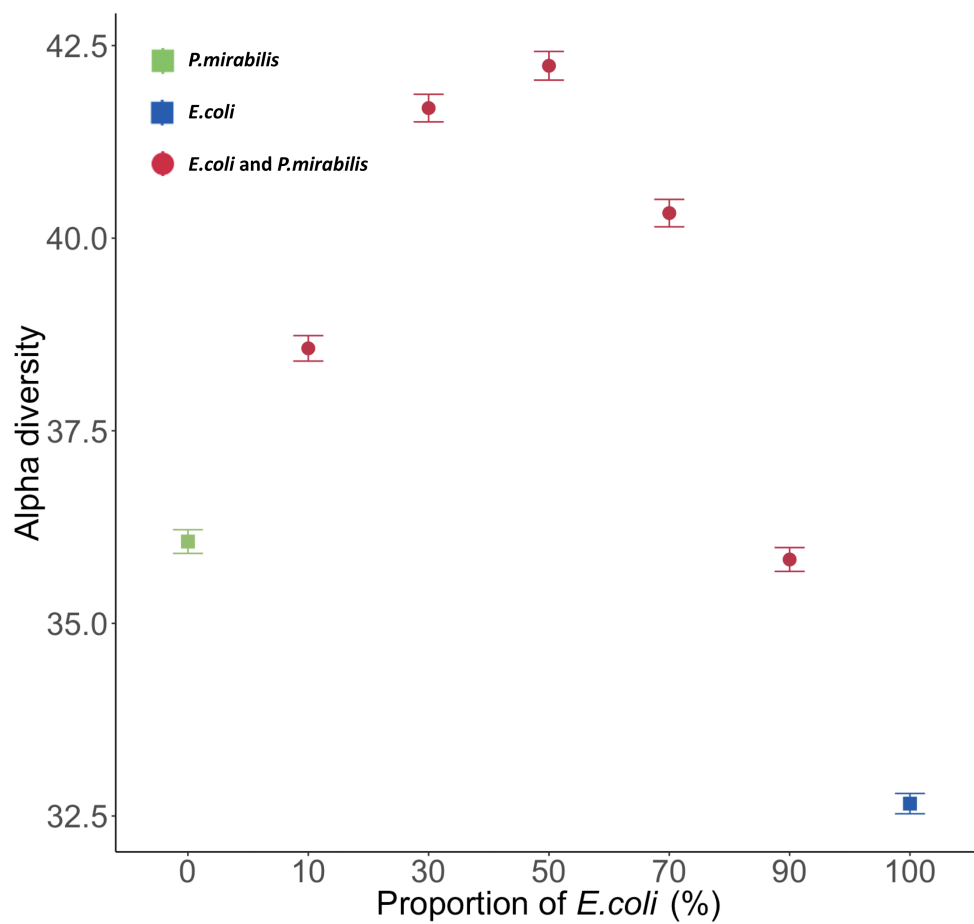


Figure S4: Within-sample diversity quantified by Nubeam using simulated samples that mixing reads at different proportions from *E. coli* and *P. mirabilis*. The bars on each dot denote  $\pm$  one standard error.

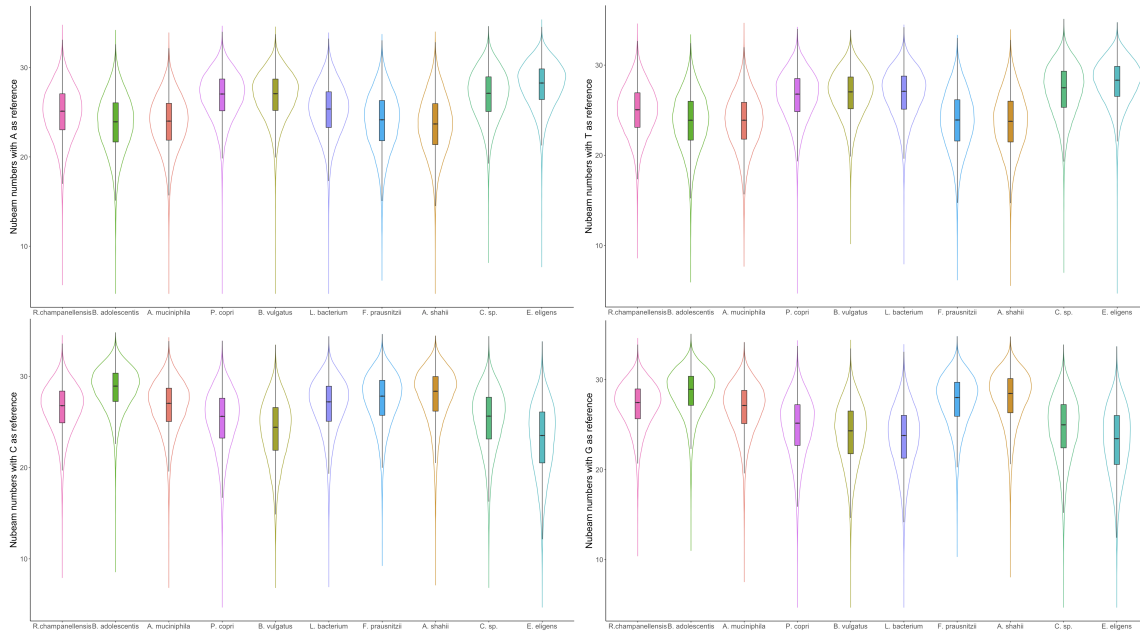
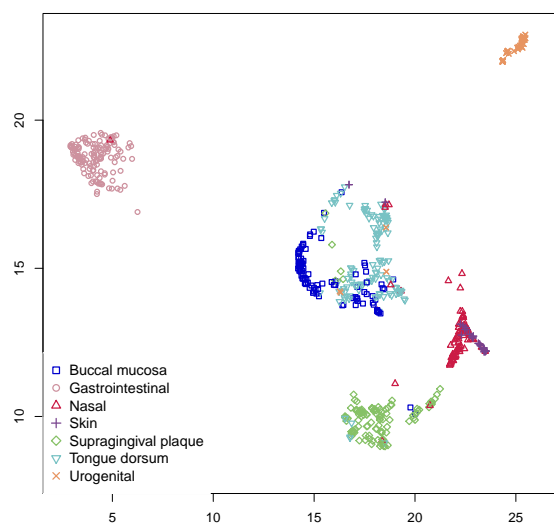
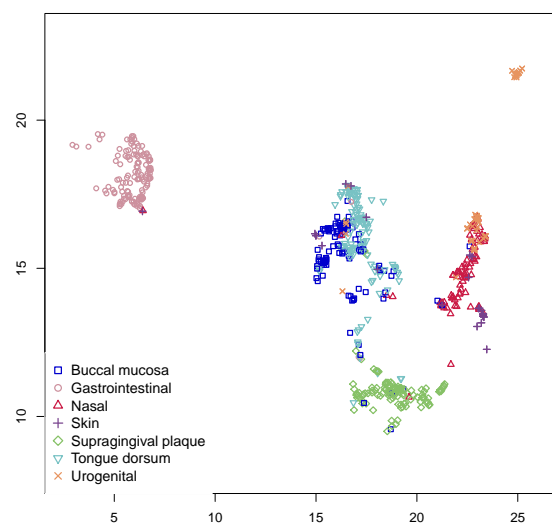


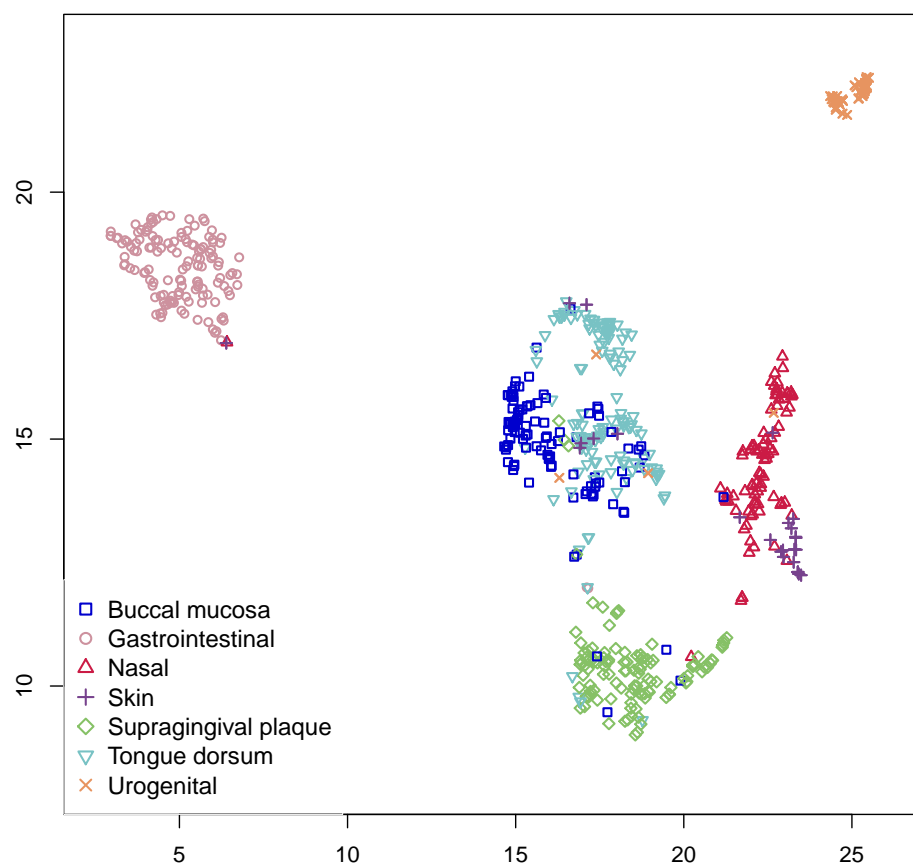
Figure S5: The distributions of Nubeam numbers for the 10 species. The species are ordered according to their individual within-sample diversity, as in Supplemental Table S3.



(a) Mapped pseudo-samples



(b) Unmapped pseudo-samples



(c) Wholesome samples

Figure S6: Samples clustering for mapped pseudo-samples, unmapped pseudo-samples, and wholesome samples respectively by UMAP.

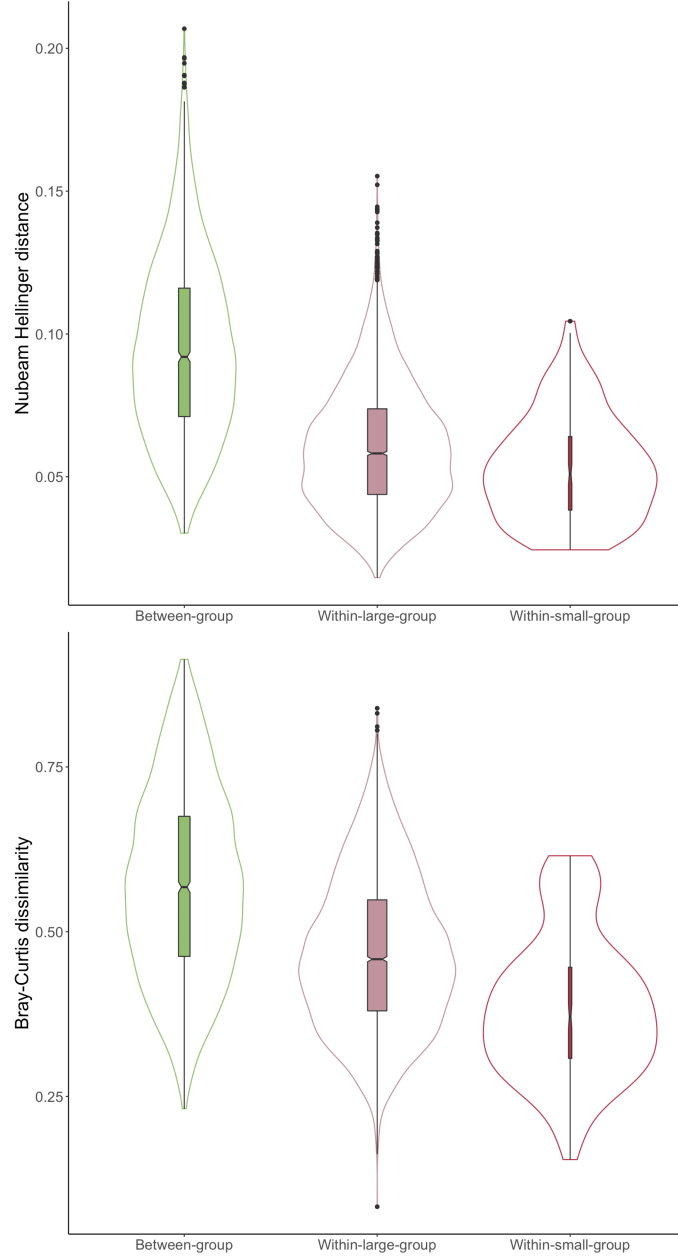


Figure S7: In Figure 6a of the main text, there are two groups of supragingival plaque mapped pseudo-samples, designated here as large and small groups, with the small group close to nasal samples. Both Nubeam and reference-based analyses showed significant difference between the compositions of the two groups: there exists significant differences among two within-group Nubeam distances and one between-group Nubeam distances (top, Kruskal-Wallis test  $p$ -value of  $5 \times 10^{-329}$ ); we quantified beta-diversity for mapped pseudo-samples using Bray-Curtis dissimilarity based on genus abundance from HMP project, there is also significant difference among two within-group dissimilarities and one between-group dissimilarities (bottom, Kruskal-Wallis test  $p$ -value of  $3 \times 10^{-154}$ ).

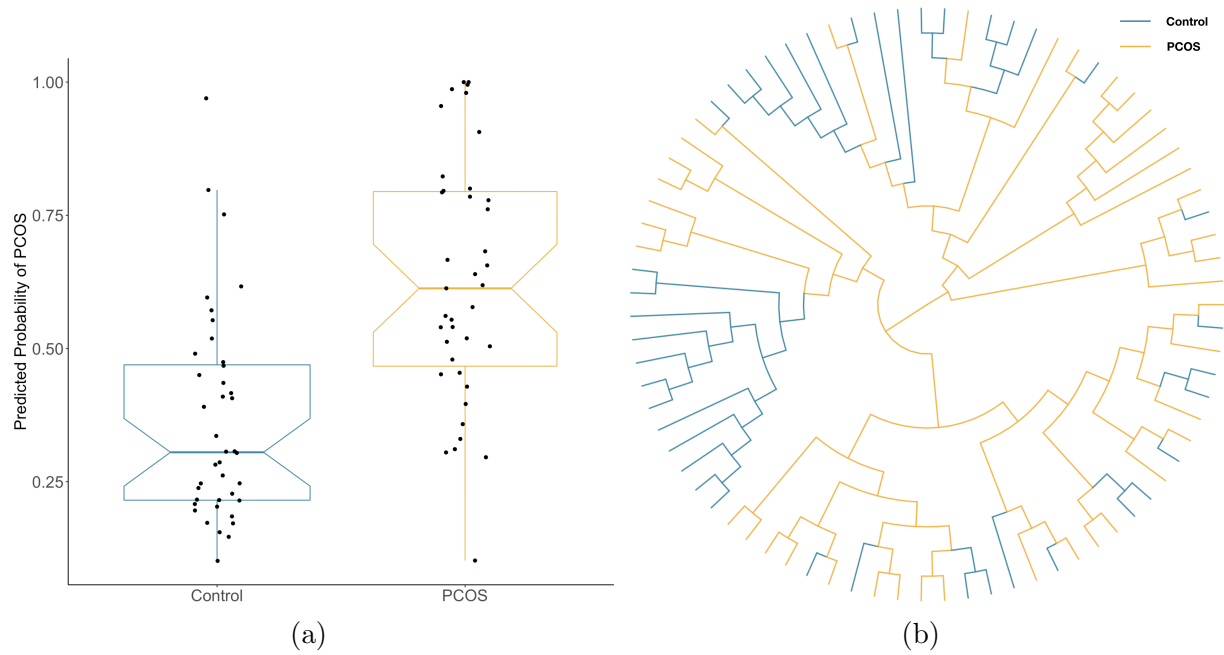


Figure S8: a: MDS was applied on Nubeam distance matrix; a logistic regression model based on the resulted first six principal coordinates was built; the predicted probability of the sample being collected from individual with PCOS for the two groups are significant (Kruskal-Wallis test  $p$ -value =  $9.443 \times 10^{-7}$ ). b: Hierarchical clustering (Wards minimum variance method) of samples based on Nubeam distance matrix.



### 3 Supplemental Tables

Tools	CPU time (s)	Memory footprint (GB)
Nubeam	10,806	27.9
CAFE 14-mer	55,388	20.7

Table S1: Runtime and memory footprint after applying Nubeam and CAFE Lu et al. (2017b) (based on 14-mer frequency) on 15 simulated WGS metagenomic samples; each sample has 50 million 75bp reads. Our benchmarking computer has two 2.3 GHz Intel Xeon E5-2699 v3 CPUs, 256 GB RAM, and running Red Hat Enterprise Linux Server 7.0.

Component	Phylum	Class	Genus
<i>R. champanellensis</i>	Firmicutes	Clostridia	<i>Ruminococcus</i>
<i>B. adolescentis</i>	Actinobacteria	Actinobacteria	<i>Bifidobacterium</i>
<i>A. muciniphila</i>	Verrucomicrobia	Verrucomicrobiae	<i>Akkermansia</i>
<i>P. copri</i>	Bacteroidetes	Bacteroidia	<i>Prevotella</i>
<i>B. vulgatus</i>	Bacteroidetes	Bacteroidia	<i>Bacteroides</i>
<i>L. bacterium</i>	Firmicutes	Clostridia	<i>Lachnospiraceae</i>
<i>F. prausnitzii</i>	Firmicutes	Clostridia	<i>Faecalibacterium</i>
<i>A. shahii</i>	Bacteroidetes	Bacteroidia	<i>Alistipes</i>
<i>C. sp.</i>	Firmicutes	Clostridia	<i>Coproccoccus</i>
<i>E. eligens</i>	Firmicutes	Clostridia	<i>Eubacterium</i>

Table S2: 10 components for synthetic communities.

Community identifier	<i>R. champanellensis</i>	<i>B. adolescentis</i>	<i>A. muciniphila</i>	<i>P. copri</i>	<i>B. vulgatus</i>	<i>L. bacterium</i>	<i>F. prausnitzii</i>	<i>A. shahii</i>	<i>C. sp.</i>	<i>E. eligens</i>	Nubeam	Simpson	Shannon
1.1	1	0	0	0	0	0	0	0	0	0	31.519	1	0
1.2	0	1	0	0	0	0	0	0	0	0	33.263	1	0
1.3	0	0	1	0	0	0	0	0	0	0	35.954	1	0
1.4	0	0	0	1	0	0	0	0	0	0	38.394	1	0
1.5	0	0	0	0	1	0	0	0	0	0	38.536	1	0
1.6	0	0	0	0	0	1	0	0	0	0	38.859	1	0
1.7	0	0	0	0	0	0	1	0	0	0	39.563	1	0
1.8	0	0	0	0	0	0	0	1	0	0	41.533	1	0
1.9	0	0	0	0	0	0	0	0	1	0	41.967	1	0
1.10	0	0	0	0	0	0	0	0	0	1	43.468	1	0
2.1	$\frac{1}{2^1}$	$\frac{1}{2^2}$	$\frac{1}{2^3}$	$\frac{1}{2^4}$	$\frac{1}{2^5}$	$\frac{1}{2^6}$	$\frac{1}{2^7}$	$\frac{1}{2^8}$	$\frac{1}{2^9}$	$\frac{1}{2^{10}}$	$38.943 \pm 0.016$	3	0.6
2.2	$\frac{1}{2^2}$	$\frac{1}{2^1}$	$\frac{1}{2^2}$	$\frac{1}{2^3}$	$\frac{1}{2^4}$	$\frac{1}{2^5}$	$\frac{1}{2^6}$	$\frac{1}{2^7}$	$\frac{1}{2^8}$	$\frac{1}{2^9}$	$45.747 \pm 0.011$	3	0.6
2.3	$\frac{1}{2^3}$	$\frac{1}{2^2}$	$\frac{1}{2^1}$	$\frac{1}{2^2}$	$\frac{1}{2^3}$	$\frac{1}{2^4}$	$\frac{1}{2^5}$	$\frac{1}{2^6}$	$\frac{1}{2^7}$	$\frac{1}{2^8}$	$44.931 \pm 0.013$	3	0.6
2.4	$\frac{1}{2^4}$	$\frac{1}{2^3}$	$\frac{1}{2^2}$	$\frac{1}{2^1}$	$\frac{1}{2^2}$	$\frac{1}{2^3}$	$\frac{1}{2^4}$	$\frac{1}{2^5}$	$\frac{1}{2^6}$	$\frac{1}{2^7}$	$42.614 \pm 0.009$	3	0.6
2.5	$\frac{1}{2^5}$	$\frac{1}{2^4}$	$\frac{1}{2^3}$	$\frac{1}{2^2}$	$\frac{1}{2^1}$	$\frac{1}{2^2}$	$\frac{1}{2^3}$	$\frac{1}{2^4}$	$\frac{1}{2^5}$	$\frac{1}{2^6}$	$45.796 \pm 0.013$	3	0.6
2.6	$\frac{1}{2^6}$	$\frac{1}{2^5}$	$\frac{1}{2^4}$	$\frac{1}{2^3}$	$\frac{1}{2^2}$	$\frac{1}{2^1}$	$\frac{1}{2^2}$	$\frac{1}{2^3}$	$\frac{1}{2^4}$	$\frac{1}{2^5}$	$47.904 \pm 0.010$	3	0.6
2.7	$\frac{1}{2^7}$	$\frac{1}{2^6}$	$\frac{1}{2^5}$	$\frac{1}{2^4}$	$\frac{1}{2^3}$	$\frac{1}{2^2}$	$\frac{1}{2^1}$	$\frac{1}{2^2}$	$\frac{1}{2^3}$	$\frac{1}{2^4}$	$46.924 \pm 0.013$	3	0.6
2.8	$\frac{1}{2^8}$	$\frac{1}{2^7}$	$\frac{1}{2^6}$	$\frac{1}{2^5}$	$\frac{1}{2^4}$	$\frac{1}{2^3}$	$\frac{1}{2^2}$	$\frac{1}{2^1}$	$\frac{1}{2^2}$	$\frac{1}{2^3}$	$52.425 \pm 0.017$	3	0.6
2.9	$\frac{1}{2^9}$	$\frac{1}{2^8}$	$\frac{1}{2^7}$	$\frac{1}{2^6}$	$\frac{1}{2^5}$	$\frac{1}{2^4}$	$\frac{1}{2^3}$	$\frac{1}{2^2}$	$\frac{1}{2^1}$	$\frac{1}{2^2}$	$47.046 \pm 0.012$	3	0.6
2.10	$\frac{1}{2^{10}}$	$\frac{1}{2^9}$	$\frac{1}{2^8}$	$\frac{1}{2^7}$	$\frac{1}{2^6}$	$\frac{1}{2^5}$	$\frac{1}{2^4}$	$\frac{1}{2^3}$	$\frac{1}{2^2}$	$\frac{1}{2^1}$	$52.670 \pm 0.010$	3	0.6
3.1	$\frac{1}{2^2}$	$\frac{1}{2^2}$	$\frac{1}{2^2}$	$\frac{1}{2^3}$	$\frac{1}{2^4}$	$\frac{1}{2^5}$	$\frac{1}{2^6}$	$\frac{1}{2^7}$	$\frac{1}{2^8}$	$\frac{1}{2^9}$	$43.073 \pm 0.010$	4.8	0.75
3.2	$\frac{1}{2^3}$	$\frac{1}{2^3}$	$\frac{1}{2^3}$	$\frac{1}{2^4}$	$\frac{1}{2^5}$	$\frac{1}{2^6}$	$\frac{1}{2^7}$	$\frac{1}{2^8}$	$\frac{1}{2^9}$	$\frac{1}{2^{10}}$	$46.948 \pm 0.013$	4.8	0.75
3.3	$\frac{1}{2^4}$	$\frac{1}{2^4}$	$\frac{1}{2^4}$	$\frac{1}{2^5}$	$\frac{1}{2^6}$	$\frac{1}{2^7}$	$\frac{1}{2^8}$	$\frac{1}{2^9}$	$\frac{1}{2^{10}}$	$\frac{1}{2^{11}}$	$45.372 \pm 0.010$	4.8	0.75
3.4	$\frac{1}{2^5}$	$\frac{1}{2^5}$	$\frac{1}{2^5}$	$\frac{1}{2^6}$	$\frac{1}{2^7}$	$\frac{1}{2^8}$	$\frac{1}{2^9}$	$\frac{1}{2^{10}}$	$\frac{1}{2^{11}}$	$\frac{1}{2^{12}}$	$45.738 \pm 0.012$	4.8	0.75
3.5	$\frac{1}{2^6}$	$\frac{1}{2^6}$	$\frac{1}{2^6}$	$\frac{1}{2^7}$	$\frac{1}{2^8}$	$\frac{1}{2^9}$	$\frac{1}{2^{10}}$	$\frac{1}{2^{11}}$	$\frac{1}{2^{12}}$	$\frac{1}{2^{13}}$	$49.527 \pm 0.019$	4.8	0.75
3.6	$\frac{1}{2^7}$	$\frac{1}{2^7}$	$\frac{1}{2^7}$	$\frac{1}{2^8}$	$\frac{1}{2^9}$	$\frac{1}{2^{10}}$	$\frac{1}{2^{11}}$	$\frac{1}{2^{12}}$	$\frac{1}{2^{13}}$	$\frac{1}{2^{14}}$	$49.754 \pm 0.012$	4.8	0.75
3.7	$\frac{1}{2^8}$	$\frac{1}{2^8}$	$\frac{1}{2^8}$	$\frac{1}{2^9}$	$\frac{1}{2^{10}}$	$\frac{1}{2^{11}}$	$\frac{1}{2^{12}}$	$\frac{1}{2^{13}}$	$\frac{1}{2^{14}}$	$\frac{1}{2^{15}}$	$51.389 \pm 0.013$	4.8	0.75
3.8	$\frac{1}{2^9}$	$\frac{1}{2^9}$	$\frac{1}{2^9}$	$\frac{1}{2^{10}}$	$\frac{1}{2^{11}}$	$\frac{1}{2^{12}}$	$\frac{1}{2^{13}}$	$\frac{1}{2^{14}}$	$\frac{1}{2^{15}}$	$\frac{1}{2^{16}}$	$53.701 \pm 0.017$	4.8	0.75
3.9	$\frac{1}{2^{10}}$	$\frac{1}{2^{10}}$	$\frac{1}{2^{10}}$	$\frac{1}{2^{11}}$	$\frac{1}{2^{12}}$	$\frac{1}{2^{13}}$	$\frac{1}{2^{14}}$	$\frac{1}{2^{15}}$	$\frac{1}{2^{16}}$	$\frac{1}{2^{17}}$	$48.698 \pm 0.012$	4.8	0.75
3.10	$\frac{1}{2^{11}}$	$\frac{1}{2^{11}}$	$\frac{1}{2^{11}}$	$\frac{1}{2^{12}}$	$\frac{1}{2^{13}}$	$\frac{1}{2^{14}}$	$\frac{1}{2^{15}}$	$\frac{1}{2^{16}}$	$\frac{1}{2^{17}}$	$\frac{1}{2^{18}}$	$49.728 \pm 0.014$	4.8	0.75
4.1	$\frac{1}{10}$	$\frac{1}{10}$	$\frac{1}{10}$	$\frac{1}{10}$	$\frac{1}{10}$	$\frac{1}{10}$	$\frac{1}{10}$	$\frac{1}{10}$	$\frac{1}{10}$	$\frac{1}{10}$	$49.210 \pm 0.015$	10	1

Table S3: In the table, each line indicates a simulated community composition; the proportion  $p_i$  of each component in the community is listed below the component. The components are arranged according to their individual Nubeam within-sample diversities (see “communities” 1.1–1.10). Except the ‘communities’ composed of only one component, each community has its  $1 \times 10^7$  reads sequencing sample simulated 10 times, with proportions adjusted according to the genome lengths of the components (Jia et al., 2013). For each community, the Nubeam within-sample diversity is compared with the reference based alpha diversity measures—inverse Simpson’s index ( $\frac{1}{\sum_i p_i^2}$ ) and Shannon index ( $-\sum_i p_i \log_{10} p_i$ ). On one hand, communities with smaller reference based measures generally have smaller Nubeam within-sample diversities; in particular, among 10 pairs of communities with similar compositions (communities 2.i and 3.i), the community with smaller reference based measures in a pair has smaller Nubeam within-sample diversity for all but one pair (2.10 and 3.10). On the other hand, communities with same reference based measures but different compositions can have different Nubeam within-sample diversities.

Body habitats	Correlation coefficient of Mantel's test	<i>p</i> -value
All	0.6883	$< 1 \times 10^{-4}$
Buccal mucosa	0.4473	$< 1 \times 10^{-4}$
Gastrointestinal	0.3652	$9 \times 10^{-4}$
Nasal	0.3805	$< 1 \times 10^{-4}$
Skin	0.4297	$8 \times 10^{-4}$
Supragingival plaque	0.6908	$< 1 \times 10^{-4}$
Tongue dorsum	0.6224	$< 1 \times 10^{-4}$
Urogenital	0.2032	$2.44 \times 10^{-2}$

Table S4: Correlation between distance matrices calculated using HMP mapped and unmapped pseudo-samples. *p*-values were based on  $10^4$  permutations.

Tools	Isolated reads	Reads from known micro-organisms	Reads from unknown micro-organisms
metaSPAdes	54%	19%	27%
MEGAHIT	74%	8%	18%

Table S5: Average proportion of reads mapped to *de novo* assembled contigs for urogenital unmapped pseudo-samples.

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